

# Exercise List: Proving convergence of the Stochastic Gradient Descent and Coordinate Descent on the Ridge Regression Problem.

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## Introduction

Consider the task of learning a rule that maps the *feature vector*  $x \in \mathbb{R}^d$  to outputs  $y \in \mathbb{R}$ . Furthermore you are given a set of labelled observations  $(x_i, y_i)$  for  $i = 1, \dots, n$ . We restrict ourselves to linear mappings. That is, we need to find  $w \in \mathbb{R}^d$  such that

$$x_i^\top w \approx y_i, \quad \text{for } i = 1, \dots, n. \quad (1)$$

That is the *hypothesis function* is parametrized by  $w$  and is given by  $h_w : x \mapsto w^\top x$ .<sup>1</sup> To choose a  $w$  such that each  $x_i^\top w$  is close to  $y_i$ , we use the squared loss  $\ell(y) = y^2/2$  and the squared regularizer. That is, we minimize

$$w^* = \arg \min_w \frac{1}{n} \sum_{i=1}^n \frac{1}{2} (x_i^\top w - y_i)^2 + \frac{\lambda}{2} \|w\|_2^2, \quad (2)$$

where  $\lambda > 0$  is the regularization parameter. We now have a complete training problem (2)<sup>2</sup>.

Using the matrix notation

$$X \stackrel{\text{def}}{=} [x_1, \dots, x_n] \in \mathbb{R}^{d \times n}, \quad \text{and} \quad y = [y_1, \dots, y_n] \in \mathbb{R}^n, \quad (3)$$

we can re-write the objective function in (2) as

$$f(w) \stackrel{\text{def}}{=} \frac{1}{2n} \|X^\top w - y\|_2^2 + \frac{\lambda}{2} \|w\|_2^2. \quad (4)$$

First we introduce some necessary notation.

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<sup>1</sup>We need only consider a linear mapping as opposed to the more general *affine* mapping  $x_i \mapsto w^\top x_i + \beta$ , because the zero order term  $\beta \in \mathbb{R}$  can be incorporated by defining a new feature vectors  $\hat{x}_i = [x_i, 1]$  and new variable  $\hat{w} = [w, \beta]$  so that  $\hat{x}_i^\top \hat{w} = x_i^\top w + \beta$

<sup>2</sup>Excluding the issue of selection  $\lambda$  using something like crossvalidation [https://en.wikipedia.org/wiki/Cross-validation\\_\(statistics\)](https://en.wikipedia.org/wiki/Cross-validation_(statistics))

**Notation:** For every  $x, w, \in \mathbb{R}^d$  let  $\langle x, w \rangle \stackrel{\text{def}}{=} x^\top w$  and let  $\|x\|_2 = \sqrt{\langle x, x \rangle}$ . Let  $A \in \mathbb{R}^{d \times d}$  be a matrix and let  $\sigma_{\min}(A)$  and  $\sigma_{\max}(A)$  be the smallest and largest singular values of  $A$  defined by

$$\sigma_{\min}(A) \stackrel{\text{def}}{=} \min_{x \in \mathbb{R}^d, x \neq 0} \frac{\|Ax\|_2}{\|x\|_2} \quad \text{and} \quad \sigma_{\max}(A) \stackrel{\text{def}}{=} \max_{x \in \mathbb{R}^d, x \neq 0} \frac{\|Ax\|_2}{\|x\|_2}. \quad (5)$$

Finally, a result you will need, if  $A$  is a symmetric positive semi-definite matrix the largest singular value of  $A$  can be defined instead as

$$\sigma_{\max}(A) = \max_{x \in \mathbb{R}^d, x \neq 0} \frac{\langle Ax, x \rangle_2}{\|x\|_2^2} = \max_{x \in \mathbb{R}^d, x \neq 0} \frac{\|Ax\|_2}{\|x\|_2}. \quad (6)$$

Therefore

$$\frac{\langle Ax, x \rangle}{\|x\|_2^2} \leq \sigma_{\max}(A), \quad \forall x \in \mathbb{R}^d \setminus \{0\}. \quad (7)$$

and

$$\frac{\|Ax\|_2}{\|x\|_2} \leq \sigma_{\max}(A), \quad \forall x \in \mathbb{R}^d \setminus \{0\}. \quad (8)$$

We will now solve the following ridge regression problem

$$w^* = \arg \min_{w \in \mathbb{R}^d} \left( \frac{1}{2n} \|X^\top w - y\|_2^2 + \frac{\lambda}{2} \|w\|_2^2 \stackrel{\text{def}}{=} f(w) \right), \quad (9)$$

using stochastic gradient descent and stochastic coordinate descent.

## Exercise 1 : Stochastic Gradient Descent (SGD)

**Some more notation:** Let  $\|A\|_F^2 \stackrel{\text{def}}{=} \text{Tr}(A^\top A)$  denote the Frobenius norm of  $A$ . Let

$$A \stackrel{\text{def}}{=} \frac{1}{n} X X^\top + \lambda I \in \mathbb{R}^{d \times d} \quad \text{and} \quad b \stackrel{\text{def}}{=} \frac{1}{n} X y. \quad (10)$$

We can exploit the separability of the objective function (2) to design a *stochastic* gradient method. For this, first we re-write the problem  $Aw = b$  as a linear least squares problem

$$\hat{w}^* = \arg \min_w \frac{1}{2} \|Aw - b\|_2^2 = \arg \min_w \sum_{i=1}^d \frac{1}{2} (A_i \cdot w - b_i)^2 \stackrel{\text{def}}{=} \arg \min_w \sum_{i=1}^d p_i f_i(w), \quad (11)$$

where  $f_i(w) = \frac{1}{2p_i} (A_i \cdot w - b_i)^2$ ,  $A_i$  denotes the  $i$ th row of  $A$ ,  $b_i$  denotes the  $i$ th element of  $b$  and  $p_i = \frac{\|A_i\|_2^2}{\|A\|_F^2}$  for  $i = 1, \dots, d$ . Note that  $\sum_{i=1}^d p_i = 1$  thus the  $p_i$ 's are probabilities.

From a given  $w^0 \in \mathbb{R}^d$ , consider the iterates

$$w^{t+1} = w^t - \alpha \nabla f_j(w^t), \quad (12)$$

where

$$\alpha = \frac{1}{\|A\|_F^2}, \quad (13)$$

and  $j$  is a random index chosen from  $\{1, \dots, d\}$  sampled with probability  $p_j$ . In other words,  $\mathbb{P}(j = i) = p_i = \frac{\|A_{i:}\|_2^2}{\|A\|_F^2}$  for all  $i \in \{1, \dots, d\}$ .

**Ex. 1** — Show that the solution  $\hat{w}^*$  to (11) and the solution to  $w^*$  to (9) are equal.

**Answer (Ex. 1)** — On the one hand, taking the gradient with respect to  $w$  in (9) and setting to zero gives

$$\frac{1}{n} X(X^\top w - y) + \lambda w = \left( \frac{1}{n} X X^\top + \lambda I \right) w - \frac{1}{n} X y \stackrel{(10)}{=} Aw - b = 0 .$$

On the other hand, differentiating (11) and setting to zero gives  $A^\top(Aw - b) = A(Aw - b) = 0$ . Since  $A$  is symmetric positive definite it is also invertible, thus

$$A(Aw - b) = 0 \Leftrightarrow (Aw - b) = 0. \quad \blacksquare$$

**Ex. 2** — Show that

$$\nabla f_j(w) = \frac{1}{p_j} A_{j:}^\top A_{j:} (w - w^*) \quad (14)$$

and that

$$\mathbb{E}_{j \sim p} [\nabla f_j(w)] \stackrel{\text{def}}{=} \sum_{i=1}^d p_i \nabla f_i(w) = A^\top A (w - w^*) ,$$

thus  $\nabla f_j(w)$  is an unbiased estimator of the full gradient of the objective function in (11). This justifies applying the stochastic gradient method.

**Answer (Ex. 2)** — First note that

$$\nabla f_j(w) = \frac{1}{p_i} A_{j:}^\top (A_{j:} w - b_j) = \frac{1}{p_i} A_{j:}^\top A_{j:} (w - w^*) .$$

Taking expectation we have that

$$\mathbb{E} [\nabla f_j(w)] = \sum_{i=1}^n \frac{p_i}{p_i} A_{j:}^\top (A_{j:} w - b_j) = A^\top (Aw - b) = A^\top A (w - w^*) . \quad \blacksquare$$

**Ex. 3** — Let  $\Pi_j \stackrel{\text{def}}{=} \frac{A_j^\top A_j}{\|A_j\|_2^2}$ , show that

$$\Pi_j \Pi_j = \Pi_j \quad , \quad (15)$$

and

$$(I - \Pi_j)(I - \Pi_j) = I - \Pi_j. \quad (16)$$

In other words,  $\Pi_j$  is a projection operator which projects orthogonally onto  $\mathbf{Range}(A_j)$ . Furthermore, if  $j \sim p_j$  verify that

$$\mathbb{E}[\Pi_j] = \sum_{i=1}^d p_i \Pi_i = \frac{A^\top A}{\|A\|_F^2}. \quad (17)$$

**Answer (Ex. 3)** — We can see that  $\Pi_j$  is an orthogonal projector by verifying that

$$\Pi_j \Pi_j = \frac{A_j^\top A_j A_j^\top A_j}{\|A_j\|_2^2 \|A_j\|_2^2} = \frac{A_j^\top \|A_j\|_2^2 A_j}{\|A_j\|_2^2 \|A_j\|_2^2} = \frac{A_j^\top A_j}{\|A_j\|_2^2} = \Pi_j \quad ,$$

and

$$(I - \Pi_j)(I - \Pi_j) = I - 2\Pi_j + \Pi_j \Pi_j \stackrel{(15)}{=} I - \Pi_j \quad .$$

Finally, we have

$$\mathbb{E}[\Pi_j] = \sum_{i=1}^m \mathbb{P}(j = i) \Pi_i = \sum_{i=1}^m \frac{\|A_i\|_2^2}{\|A\|_F^2} \frac{A_i^\top A_i}{\|A_i\|_2^2} = \sum_{i=1}^m \frac{A_i^\top A_i}{\|A\|_F^2} = \frac{A^\top A}{\|A\|_F^2}. \quad \blacksquare$$

**Ex. 4** — Show that the distance to the solution satisfies the following recurrence

$$\|w^{t+1} - w^*\|_2^2 = \|w^t - w^*\|_2^2 - \left\langle \frac{A_j^\top A_j}{\|A_j\|_2^2} (w^t - w^*), w^t - w^* \right\rangle. \quad (18)$$

**Answer (Ex. 4)** — Using (14) and subtracting  $w^*$  from both sides of (12) we have

$$\begin{aligned} w^{t+1} - w^* &= w^t - w^* - \frac{\alpha_j}{p_j} A_j^\top A_j (w^t - w^*) \\ &\stackrel{(13)}{=} \left( I - \frac{A_j^\top A_j}{\|A_j\|_2^2} \right) (w^t - w^*). \end{aligned}$$

Taking norm squared in the above we have that

$$\begin{aligned}
\|w^{t+1} - w^*\|_2^2 &= \left\| \left( I - \frac{A_j^\top A_j}{\|A_j\|_2^2} \right) (w^t - w^*) \right\|_2^2 \\
&\stackrel{(15)}{=} \left\langle \left( I - \frac{A_j^\top A_j}{\|A_j\|_2^2} \right) (w^t - w^*), w^t - w^* \right\rangle \\
&= \|w^t - w^*\|_2^2 - \left\langle \frac{A_j^\top A_j}{\|A_j\|_2^2} (w^t - w^*), w^t - w^* \right\rangle.
\end{aligned}$$

**Ex. 5** — Using the solutions to the previous equations show that the iterates (12) converge according to

$$\mathbb{E} [\|w^{t+1} - w^*\|_2^2] \leq \left( 1 - \frac{\sigma_{\min}(A)^2}{\|A\|_F^2} \right) \mathbb{E} [\|w^t - w^*\|_2^2]. \quad (19)$$

**Answer (Ex. 5)** — Taking expectation conditioned on  $w^t$  in the above gives

$$\begin{aligned}
\mathbb{E} [\|w^{t+1} - w^*\|_2^2 | w^t] &= \|w^t - w^*\|_2^2 - \left\langle \mathbb{E} \left[ \frac{A_j^\top A_j}{\|A_j\|_2^2} \right] (w^t - w^*), w^t - w^* \right\rangle \\
&\stackrel{(17)}{=} \|w^t - w^*\|_2^2 - \frac{1}{\|A\|_F^2} \left\langle A^\top A (w^t - w^*), w^t - w^* \right\rangle \\
&\stackrel{(7)}{\leq} \|w^t - w^*\|_2^2 - \frac{\sigma_{\min}(A)^2}{\|A\|_F^2} \|w^t - w^*\|_2^2 \\
&= \left( 1 - \frac{\sigma_{\min}(A)^2}{\|A\|_F^2} \right) \|w^t - w^*\|_2^2.
\end{aligned}$$

It remains to take expectation in the above. ■

**Remark:** This is an amazing and recent result [2], since it shows that SGD converges exponentially fast despite the fact that the iterates (14) only require access to a single row of  $A$  at a time! This result can be extended to solving any linear system  $Aw = b$ , including the case where  $A$  rank deficient. Indeed, so long as there exists a solution to  $Aw = b$ , the iterates (14) converge to the solution of least norm and at rate of  $\left( 1 - \frac{\sigma_{\min}^+(A)^2}{\|A\|_F^2} \right)$  where  $\sigma_{\min}^+(A)$  is the smallest nonzero singular value of  $A$  [1]. Thus this method can solve any linear system.

## BONUS

### Exercise 2: Stochastic Coordinate Descent (CD)

Consider the minimization problem

$$w^* = \arg \min_{x \in \mathbb{R}^d} \left( f(w) \stackrel{\text{def}}{=} \frac{1}{2} w^\top A w - w^\top b \right), \quad (20)$$

where  $A \in \mathbb{R}^{d \times d}$  is a symmetric positive definite matrix, and  $w, b \in \mathbb{R}^d$ .

**Ex. 6** — First show that, using the notation (10), solving (20) is equivalent to solving (9).

**Answer (Ex. 6)** — Differentiating (20) or (9) in  $w$  gives

$$\nabla f(x) = Ax - b.$$

Consequently the unique solution  $w^*$  to both of these problems is given by  $w^* = A^{-1}b$ . ■

**Ex. 7** — Show that

$$\frac{\partial f(w)}{\partial w_i} = A_{i:} w - b_i, \quad (21)$$

where  $A_{i:}$  is the  $i$ th row of  $A$ . Furthermore note that  $w^* = A^{-1}b$ , thus

$$\frac{\partial f(w)}{\partial w_i} = e_i^\top (Aw - b) = e_i^\top A(w - w^*). \quad (22)$$

**Answer (Ex. 7)** — Follows immediately from  $\nabla f(x) = Ax - b$  and  $w^* = A^{-1}b$ . ■

**Ex. 8** — **Question 2.3:** Consider a step of the stochastic coordinate descent method

$$w^{k+1} = w^k - \alpha_i \frac{\partial f(w^k)}{\partial x_i} e_i, \quad (23)$$

where  $e_i \in \mathbb{R}^d$  is the  $i$ th unit coordinate vector,  $\alpha_i = \frac{1}{A_{ii}}$ , and  $i \in \{1, \dots, d\}$  is sampled i.i.d at each step according to  $i \sim p_i$  where  $p_i = \frac{A_{ii}}{\text{Tr}(A)}$ . Let  $\|x\|_A^2 \stackrel{\text{def}}{=} x^\top A x$ .

First, let  $\Pi_i = \frac{e_i e_i^\top A}{A_{ii}}$  and prove that

$$\|w^{k+1} - w^*\|_A^2 = \left\langle (I - \Pi_i^\top) A (I - \Pi_i) (w^k - w^*), w^k - w^* \right\rangle. \quad (24)$$

**Answer (Ex. 8)** — Subtracting  $w^*$  from both sides of (23) gives

$$\begin{aligned} w^{k+1} - w^* &\stackrel{(22)+(23)}{=} w^k - w^* - \alpha_i e_i^\top A (w^k - w^*) e_i \\ &= \left( I - \frac{e_i e_i^\top A}{A_{ii}} \right) (w^k - w^*). \end{aligned} \quad (25)$$

Taking the squared norm  $\|\cdot\|_A$  on both sides of (25) gives

$$\begin{aligned} \|w^{k+1} - w^*\|_A^2 &= \left\langle A(I - \Pi_i)(w^k - w^*), (I - \Pi_i)(w^k - w^*) \right\rangle \\ &= \left\langle (I - \Pi_i^\top)A(I - \Pi_i)(w^k - w^*), w^k - w^* \right\rangle. \end{aligned}$$

■

**Ex. 9 — Question 2.4:** Let  $r^k \stackrel{\text{def}}{=} A^{1/2}(w^k - w^*)$ . Deduce from (24) that

$$\|r^{k+1}\|_2^2 = \|r^k\|_2^2 - \left\langle \frac{A^{1/2} e_i e_i^\top A^{1/2}}{A_{ii}} r^k, r^k \right\rangle. \quad (26)$$

**Answer (Ex. 9)** — Let  $r^k = A^{1/2}(w^k - w^*)$  and note that

$$(I - \Pi_i^\top)A(I - \Pi_i) = A - 2A\Pi_i + \Pi_i^\top A\Pi_i = A - \frac{Ae_i e_i^\top A}{A_{ii}}.$$

Using this we have from (24) that

$$\begin{aligned} \|r^{k+1}\|_2^2 &= \left\langle \left( A - \frac{Ae_i e_i^\top A}{A_{ii}} \right) (w^k - w^*), w^k - w^* \right\rangle \\ &= \|r^k\|_2^2 - \left\langle \frac{Ae_i e_i^\top A}{A_{ii}} (w^k - w^*), w^k - w^* \right\rangle \\ &= \|r^k\|_2^2 - \left\langle \frac{A^{1/2} e_i e_i^\top A^{1/2}}{A_{ii}} r^k, r^k \right\rangle. \quad \blacksquare \end{aligned} \quad (27)$$

**Ex. 10** — Finally, prove the convergence of the iterates of CD (23) converge according to

$$\mathbb{E} \left[ \|w^{k+1} - w^*\|_A^2 \right] \leq \left( 1 - \frac{\lambda_{\min}(A)}{\text{Tr}(A)} \right) \mathbb{E} \left[ \|w^k - w^*\|_A^2 \right] \quad (28)$$

thus (23) converges to the solution.

**Hint:** Since  $A$  is symmetric positive definite you can use that

$$\lambda_{\min}(A) = \inf_{x \in \mathbb{R}^d, x \neq 0} \frac{x^\top A x}{\|x\|_2^2}.$$

You will need to use that  $x^\top A x \geq \lambda_{\min}(A)\|x\|_2^2$  at some point.

**Answer (Ex. 10)** — Taking expectation conditioned on  $r^k$  over the second term in (27) gives

$$\begin{aligned} \mathbb{E} \left[ \left\langle \frac{A^{1/2} e_i e_i^\top A^{1/2}}{A_{ii}} r^k, r^k \right\rangle \mid r^k \right] &= \sum_{j=1}^n \frac{A_{jj}}{\text{Tr}(A)} \left\langle \frac{A^{1/2} e_j e_j^\top A^{1/2}}{A_{jj}} r^k, r^k \right\rangle \\ &= \frac{1}{\text{Tr}(A)} \left\langle A^{1/2} \sum_{j=1}^n e_j e_j^\top A^{1/2} r^k, r^k \right\rangle \\ &= \frac{1}{\text{Tr}(A)} \left\langle A r^k, r^k \right\rangle \\ &\geq \frac{\lambda_{\min}(A)}{\text{Tr}(A)} \|r^k\|_2^2. \end{aligned}$$

Consequently taking expectation conditioned on  $r^k$  in (26) gives

$$\mathbb{E} \left[ \|r^{k+1}\|_2^2 \mid r^k \right] \leq \left( 1 - \frac{\lambda_{\min}(A)}{\text{Tr}(A)} \right) \|r^k\|_2^2. \quad (29)$$

It now remains to take expectation and re-write  $\|r^k\|_2^2 = \|w^k - w^*\|_A^2$ . ■

**Ex. 11 — Question 2.6:** When is this stochastic coordinate descent method *faster* than the stochastic gradient method (14) or gradient descent? Note that each iteration of SGD and CD costs  $O(d)$  floating point operations while an iteration of the GD method costs  $O(d^2)$  floating point operations (assuming that  $A$  has been previously calculated and stored). What happens if  $d$  is very big? What if  $\text{Tr}(A)$  is very large? Discuss this.

**Answer (Ex. 11)** — Let

$$\kappa_{SGD} \stackrel{\text{def}}{=} \frac{\|A\|_F^2}{\sigma_{\min}^2(A)} = \frac{\text{Tr}(A^\top A)}{\sigma_{\min}^2(A)} = \sum_{i=1}^d \frac{\sigma_i^2(A)}{\sigma_{\min}^2(A)},$$



be the complexity constant of SGD and let

$$\kappa_{CD} \stackrel{\text{def}}{=} \frac{\text{Tr}(A)}{\lambda_{\min}(A)} = \sum_{i=1}^d \frac{\sigma_i(A)}{\sigma_{\min}(A)},$$

be the complexity constant of CD, where we used that  $A$  is positive semi-definite so that  $\lambda_i(A) = \sigma_i(A)$ .

Consider the extreme case where  $\sigma_i(A) = \sigma_j(A)$  for every  $i, j \in \{1, \dots, d\}$ . In this case  $\kappa_{SGD} = d = \kappa_{CD}$ .

Now consider the case that the singular values are evenly spread out with  $\sigma_i(A) = i \times \tau$  where  $\tau > 0$ . In this case

$$\kappa_{SGD} = \sum_{i=1}^d \frac{i^2 \times \tau^2}{\tau^2} = O(d^3)$$

and

$$\kappa_{CD} = \sum_{i=1}^d \frac{i \times \tau}{\tau} = O(d^2).$$

Essentially, the complexity of coordinated descent  $\kappa_{CD}$  is far less sensitive to *ill-conditioned* data, that is, data where the smallest and the largest singular values are far apart.

## References

- [1] R. M. Gower and P. Richtárik. “Stochastic Dual Ascent for Solving Linear Systems”. In: *arXiv:1512.06890* (2015).
- [2] T. Strohmer and R. Vershynin. “A Randomized Kaczmarz Algorithm with Exponential Convergence”. In: *Journal of Fourier Analysis and Applications* 15.2 (2009), pp. 262–278.